On the application of common principal components in biplots

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- **1** What are common principal components (CPCs)?
- **2** Identifying the CPCs
- **3** Simultaneous diagonalisation methods
- Application of the CPC model in biplots

5 Conclusions

What are common principal components (CPCs)?

How can variance structures of two (or more) groups differ? Univariate case:

• Homoscedastic or heteroscedastic (nothing in between)

Multivariate case:

- Number of different ways covariance matrices can differ (Flury 1988):
 - $\bullet \quad \mathsf{Equality} \ \boldsymbol{\Sigma}_1 = \boldsymbol{\Sigma}_2$
 - 2 Proportionality $\Sigma_1 = \rho \Sigma_2$
 - Ommon principal components
 - Partial common principal components
 - 6 Heterogeneity

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Flury's hierarchy: Common principal components (CPC)

Flury's hierarchy: Heterogeneity





Principal component analysis (PCA):

 $\pmb{\Sigma}=\pmb{B}\pmb{\Lambda}\pmb{B}'$

Common principal components (CPC):

 $\mathbf{\Sigma}_1 = \mathbf{B} \mathbf{\Lambda}_1 \mathbf{B}'$ $\mathbf{\Sigma}_2 = \mathbf{B} \mathbf{\Lambda}_2 \mathbf{B}'$

Partial common principal components (CPC(q)):

$$\begin{split} \boldsymbol{\Sigma}_1 &= \boldsymbol{\mathsf{B}}_1 \boldsymbol{\mathsf{\Lambda}}_1 \boldsymbol{\mathsf{B}}_1' \quad \text{where} \quad \boldsymbol{\mathsf{B}}_1 &= [\boldsymbol{\mathsf{b}}_1 \dots \boldsymbol{\mathsf{b}}_q : \boldsymbol{\mathsf{b}}_{q+1(1)} \dots \boldsymbol{\mathsf{b}}_{p(1)}] \\ \boldsymbol{\Sigma}_2 &= \boldsymbol{\mathsf{B}}_2 \boldsymbol{\mathsf{\Lambda}}_2 \boldsymbol{\mathsf{B}}_2' \qquad \qquad \boldsymbol{\mathsf{B}}_2 &= [\boldsymbol{\mathsf{b}}_1 \dots \boldsymbol{\mathsf{b}}_q : \boldsymbol{\mathsf{b}}_{q+1(2)} \dots \boldsymbol{\mathsf{b}}_{p(2)}] \end{split}$$

Advantages the CPC model might provide:

- more stable estimates than when incorrectly assuming *heterogeneity* of covariance matrices
- more accurate estimates than when incorrectly assuming equality of covariance matrices

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Identifying the CPCs

Table 7.9. Decomposition of X_{total}^2 in Head Dimension Example (k = 2, p = 6)

| Model | | X ² | df | X ² | AIC for |
|-----------------|-----------------|----------------|----|----------------|--------------|
| Higher | Lower | | | dſ | Higher Model |
| Equality | Proportionality | 42.29 | 1 | 42.29 | 89.78 |
| Proportionality | CPC | 25.66 | 5 | 5.13 | 49.49 |
| CPC | CPC(1) | 15.12 | 10 | 1.51 | 33.82* |
| CPC(1) | Unrelated | 6.70 | 5 | 1.34 | 38.70 |
| Unrelated | | | | | 42.0 |
| Equality | Unrelated | 89.78 | 21 | | |

*Minimum AIC.

- The χ² statistics are *not independent* and *assume normality* of the k populations
- The AIC is not a formal hypothesis test

Different approach (Krzanowski 1979)

Geometrically: dot product of two unit vectors \mathbf{a} and \mathbf{b} = cosine of the angle between the two vectors in *p*-dimensional space.



 Do pairwise comparisons of the dot products from all combinations of the p principal components from k groups.

Simulated CPC data, k = 2, p = 5, n = 200

• Arbitrary cut-off point: $\cos^{-1}(0.95) = 18.2$ degrees



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Simulated CPC(2) data, k = 2, p = 5, n = 200



Dot product values for the permutations

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Simultaneous diagonalisation methods

• FG algorithm (Flury 1988)

$$\min \phi(\mathbf{\Lambda}_i) := rac{\det(\operatorname{diag}(\mathbf{\Lambda}_i))}{\det(\mathbf{\Lambda}_i)}$$

- Stepwise CPC (Trendafilov 2010)
- rjd/JADE (Cardoso & Souloumiac 1996)

$$\min\sum_{i=1}^p \sum_{j>i}^p \lambda_{ij}^2$$

Compared these with:

- Eigenvectors of the pooled covariance matrix
- Eigenvectors of the covariance matrix of the pooled data

Application of the CPC model in biplots

Swiss bank notes data:



- X_1 : Length of the bank note,
- X_2 : Height of the bank note, measured on the left,
- X_3 : Height of the bank note, measured on the right,
- X_4 : Distance of inner frame to the lower border,
- X_5 : Distance of inner frame to the upper border,

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 X_6 : Length of the diagonal.

Stepwise CPC biplot: Bank notes data

Pooled covariance matrix biplot: Bank notes data



Pooled data biplot: Bank notes data



Flury CPC biplot: Bank notes data



Biplot goodness of fit

Overall quality of the display (Gower, Lubbe & Le Roux 2011)

Letting **X** contain the data from all *k* groups, with the columns of **X** centred, and $||\mathbf{X}||^2 = tr(\mathbf{X}'\mathbf{X})$, the total variation in the data can be partitioned as follows:

$$||\mathbf{X}||^2 = ||\hat{\mathbf{X}}_{[r]}||^2 + ||\mathbf{X} - \hat{\mathbf{X}}_{[r]}||^2$$

Total goodness of fit =
$$\frac{||\hat{\mathbf{X}}_{[r]}||^2}{||\mathbf{X}||^2} = \frac{\sum_{i=1}^r \lambda_i}{\sum_{i=1}^p \lambda_i}$$

Biplot goodness of fit

Within group variation

Letting X_i contain the data from the i^{th} group, with the columns of X_i centred *per group*, the quality of representation of the within group variation can be measured as follows:

Within groups goodness of fit =
$$\frac{\sum_{i=1}^{k} ||\hat{\mathbf{X}}_{[r]}||^2}{\sum_{i=1}^{k} ||\mathbf{X}||^2} = \frac{\sum_{j=1}^{k} \sum_{i=1}^{r} \lambda_{ji}}{\sum_{j=1}^{k} \sum_{i=1}^{p} \lambda_{ji}}$$

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Simulated CPC data: k = 2, p = 5, n = 200



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Pooled covariance matrix biplot: Simulated CPC data



Pooled data biplot: Simulated CPC data







Simulated CPC(2) data: k = 2, p = 5, n = 200



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Pooled covariance matrix biplot: Simulated CPC(2) data



Pooled data biplot: Simulated CPC(2) data

Flury CPC biplot: Simulated CPC(2) data



Conclusions

- Eigenvectors of the covariance matrix of the *pooled data* provide the simplest and best quality display for grouped data in 2D or 3D biplots
- Preliminary work also indicates that the axis predictivities (quality of representation of the variables) of the pooled data biplot are higher than for CPC biplots
- Eigenvectors of the pooled covariance matrix and the CPC solutions provide similar quality biplot displays
- CPC solutions are more useful for maximising the variation *within* groups than the variation *between* groups

Sources

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